

# Multi-level Monte Carlo methods in Uncertainty Quantification

Fabio Nobile

CSQI - Institute of Mathematics, EPFL, Switzerland

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EU-FP7 project: *Uncertainty Management for Robust Industrial Design in Aeronautics (UMRIDA)*



Center for Advanced Modeling Science



# Outline

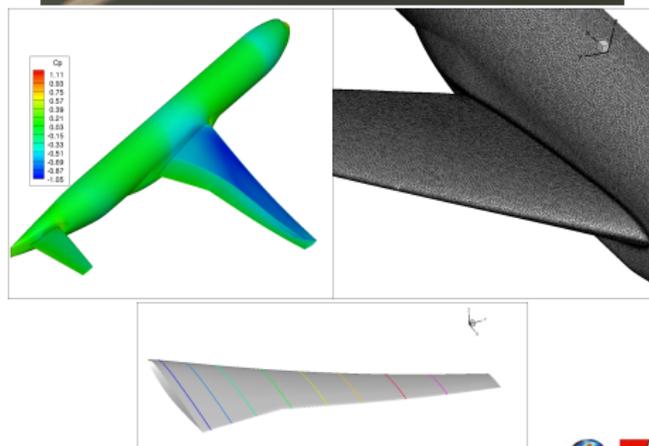
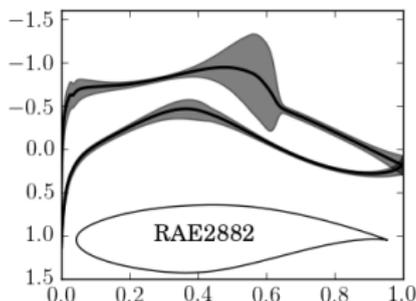
- 1 Motivating example
- 2 Multilevel Monte Carlo for expectations
- 3 MLMC for moments and distributions
- 4 Risk averse optimization with MLMC
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# UQ in aerodynamic design

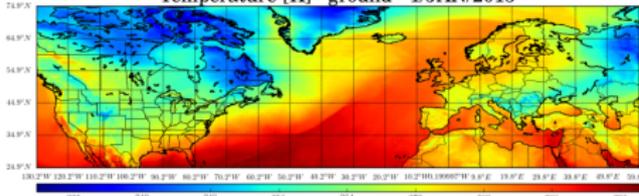
Compute aerodynamic coeffs. (lift, drag,  $C_p$ ) and optimize airfoil shape in presence of operational uncertainties (Mach number, angle of attack, ...) and geometrical uncertainties (manufacturing tolerances, icing, fatigue, ...)



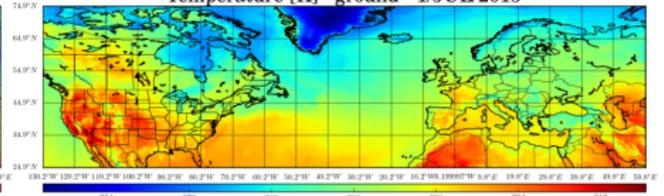
# Operational uncertainties

Atmospheric fluctuations with respect to location, time ( $T, p, \rho, \mathbf{u}$ ) over long flights

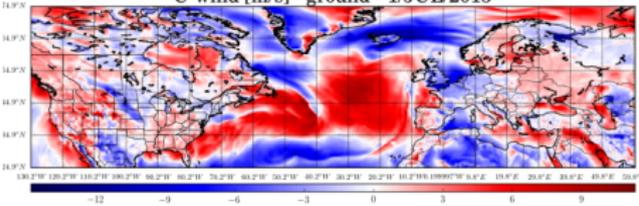
Temperature [K] - ground - 1/JAN/2015



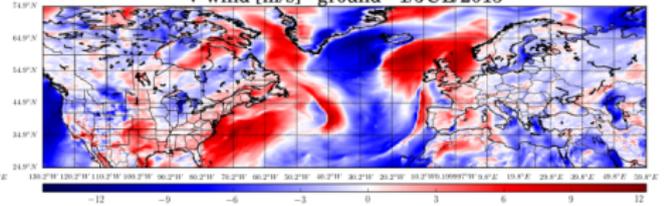
Temperature [K] - ground - 1/JUL/2015



U wind [m/s] - ground - 1/JUL/2015



V wind [m/s] - ground - 1/JUL/2015



**Probabilistic framework:** Mach, Reynolds, Angle of Attack, etc. treated as random variables

# Geometrical uncertainties

**Production:** manufacturing, assembly



**Temporary factors:** deflection, icing



**Permanent/degrading factors:** impacts, erosion, fouling



**Probabilistic Framework:** Leading edge radius, thickness, curvature, etc. treated as random variables

# Forward Uncertainty propagation

- **Random input parameters:**  $y$  (with given distribution)
- **(Complex) Model:**  $\mathcal{L}_y u = \mathcal{F}$  (e.g. Euler, Navier-Stokes,...)  
hence  $u = u(y)$  is a random solution
- **Quantity of interest:**  $Q = Q(u)$  (random output, e.g. lift, drag, etc.)

**Goal:** compute  $\mu(Q) = \mathbb{E}[Q]$  or other statistical quantities

In practice,  $u$  is not accessible. **Computational model**

$$\mathcal{L}_{h,y} u_h = \mathcal{F}_h \quad \implies \quad \text{computational output} \quad Q_h = Q(u_h)$$

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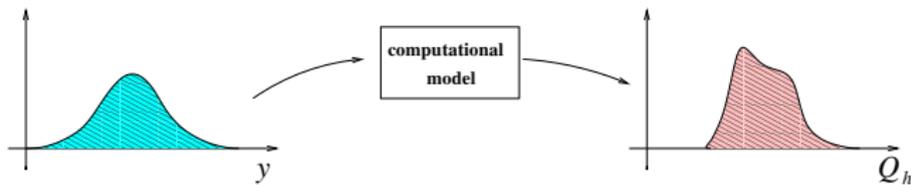
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# Monte Carlo method

- Generate  $M$  iid copies  $y^{(1)}, \dots, y^{(M)} \sim y$
- Compute the corresponding outputs  $Q_h^{(i)}$ ,  $i = 1, \dots, M$
- Approximate expectation by sample average

$$\mu_h^{MC} = \frac{1}{M} \sum_{i=1}^M Q_h^{(i)} \quad (\text{biased estimator } \mathbb{E}[\mu_h^{MC}] = \mathbb{E}[Q_h] \neq \mathbb{E}[Q])$$

## Mean squared error

$$\text{MSE}(\mu_h^{MC}) := \mathbb{E}[(\mu(Q) - \mu_h^{MC})^2] = \underbrace{(\mathbb{E}[Q - Q_h])^2}_{\text{discret. error}} + \underbrace{\frac{\text{Var}[Q_h]}{M}}_{\text{MC error}}$$

## Complexity analysis (error versus cost)

Assume: •  $|\mathbb{E}[Q - Q_h]| = \mathcal{O}(h^\alpha)$ ,  $\text{Var}[Q_h] = \mathcal{O}(1)$ ,

• cost to compute each  $Q_h^{(i)}$ :  $C_h = \mathcal{O}(h^{-\gamma})$

Then  $\text{MSE} = \mathcal{O}(\text{tol}^2) \implies h = \mathcal{O}(\text{tol}^{\frac{1}{\alpha}})$ ,  $M = \mathcal{O}(\text{tol}^{-2})$

Total work:  $\text{Work}(\mu_h^{MC}) = C_h M \lesssim \text{tol}^{-\frac{\gamma}{\alpha}} \text{tol}^{-2}$

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# Can we improve on Monte Carlo? Control variate

Let  $Z$  be random variable correlated with  $Q_h$ , and with known mean.

**Idea:** Apply MC on  $Q_{h,Z} = Q_h - \alpha(Z - \mathbb{E}[Z])$  (notice that  $\mathbb{E}[Q_{h,Z}] = \mathbb{E}[Q_h]$ )

$$\mu_h^{CV} = \frac{1}{M} \sum_{i=1}^M (Q_h^{(i)} - \alpha Z^{(i)}) + \alpha \mathbb{E}[Z]$$

$$\text{Var}[Q_{h,Z}] = \text{Var}[Q_h - \alpha Z] = \text{Var}[Q_h] + \alpha^2 \text{Var}[Z] - 2\alpha \text{Cov}(Q_h, Z)$$

For optimal  $\alpha$ :  $\text{Var}[Q_{h,Z}] = \text{Var}[Q_h] \left(1 - \frac{\text{Cov}(Q_h, Z)}{\text{Var}[Z]}\right) \leq \text{Var}[Q_h]$  (always gives variance reduction)

## Two ideas for choosing $Z$

- Use a surrogate model  $Z = Q^{surr}$  with numerically optimized  $\alpha$   
 $\rightsquigarrow$  *multi-fidelity Monte Carlo* [Peherstorfer, Willcox, Gunzburger, 2016]
- Use coarser discretization e.g.  $Z = Q_{2h}$  (usually with  $\alpha = 1$ )  
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**Problem:**  $\mathbb{E}[Z]$  not known, in general !

↪ compute it with independent MC with larger sample size (cheaper problem).

From two-level to multilevel:

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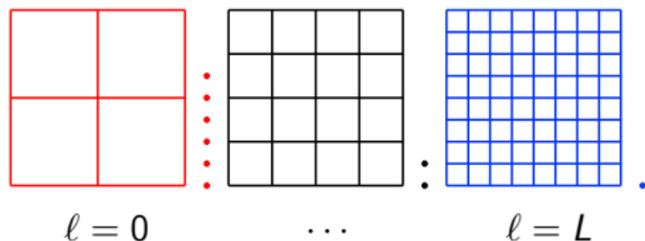
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# Multilevel Monte Carlo



- Sequence of refined discretizations

$$h_0 > h_1 > \dots > h_L$$

- Sequence of sample sizes

$$M_0 > M_1 > \dots > M_L$$

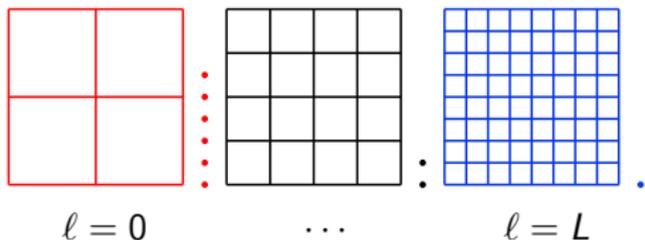
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$$\mu_L^{MLMC} = \sum_{\ell=0}^L \frac{1}{M_\ell} \sum_{i=1}^{M_\ell} (Q_\ell^{(i,\ell)} - Q_{\ell-1}^{(i,\ell)}), \quad Q_{-1} = 0$$

## Mean squared error

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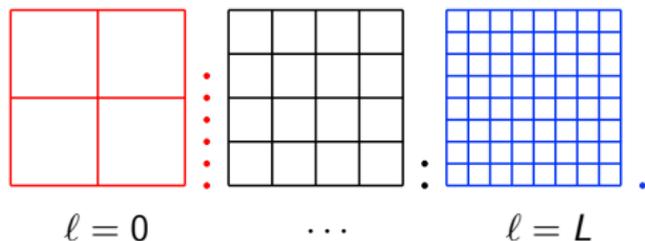
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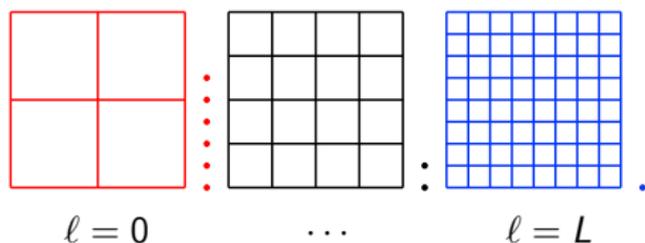
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# Multilevel Monte Carlo

- $V_\ell = \text{Var}[Q_\ell - Q_{\ell-1}]$  (variance of differences)
- $C_\ell = \text{cost of computing each } \Delta Q_\ell^{(i,\ell)} = Q_\ell^{(i,\ell)} - Q_{\ell-1}^{(i,\ell)}$

**Optimal sample sizes  $M_\ell$ :** [Giles 2008] minimize  $W = \sum_{\ell=0}^L C_\ell M_\ell$  s.t.  $\text{MSE} \simeq \text{tol}^2$

$$M_\ell = \left\lceil \text{tol}^{-2} \sqrt{\frac{V_\ell}{C_\ell}} \left( \sum_{k=0}^L \sqrt{C_k V_k} \right) \right\rceil$$

**Complexity analysis for  $h_\ell = h_0 s^{-\ell}$ :** [Giles 2008, Cliffe-Giles-Scheichl-Teckentrup 2011]

Assume

- $|\mathbb{E}[Q - Q_\ell]| = \mathcal{O}(h_\ell^\alpha)$ ,
- $V_\ell = \text{Var}[Q_\ell - Q_{\ell-1}] = \mathcal{O}(h_\ell^\beta)$ , ( $\beta = 2\alpha$  for smooth problems/noise)
- $C_\ell = \mathcal{O}(h_\ell^{-\gamma})$ ,  $2\alpha \geq \min\{\beta, \gamma\}$

Then, choosing  $L = \mathcal{O}(\text{tol}^{\frac{1}{\alpha}})$  and  $M_\ell$  as above gives  $\text{MSE}(\mu_L^{\text{MLMC}}) \leq \text{tol}^2$  and

$$\text{Work}(\mu_L^{\text{MLMC}}) = \sum_{\ell=0}^L C_\ell M_\ell \lesssim \begin{cases} \text{tol}^{-2}, & \beta > \gamma \\ \text{tol}^{-2} (\log \text{tol})^2, & \beta = \gamma \\ \text{tol}^{-2 - \frac{\gamma - \beta}{\alpha}}, & \beta < \gamma \end{cases}$$

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- $V_\ell = \text{Var}[Q_\ell - Q_{\ell-1}] = \mathcal{O}(h_\ell^\beta)$ , ( $\beta = 2\alpha$  for smooth problems/noise)
- $C_\ell = \mathcal{O}(h_\ell^{-\gamma})$ ,  $2\alpha \geq \min\{\beta, \gamma\}$

Then, choosing  $L = \mathcal{O}(\text{tol}^{\frac{1}{\alpha}})$  and  $M_\ell$  as above gives  $\text{MSE}(\mu_L^{\text{MLMC}}) \leq \text{tol}^2$  and

$$\text{Work}(\mu_L^{\text{MLMC}}) = \sum_{\ell=0}^L C_\ell M_\ell \lesssim \begin{cases} \text{tol}^{-2}, & \beta > \gamma \\ \text{tol}^{-2} (\log \text{tol})^2, & \beta = \gamma \\ \text{tol}^{-2 - \frac{\gamma - \beta}{\alpha}}, & \beta < \gamma \end{cases}$$

# Multilevel Monte Carlo – practical aspects

**Remark:** MC complexity always improved for optimal choice of  $M_\ell$ .  
 For  $\beta = 2\alpha$  we get either  $\mathcal{O}(\text{tol}^{-2})$  (up to log terms) or  $\mathcal{O}(\text{tol}^{-\frac{\gamma}{\alpha}})$ .

To achieve improved complexity, one needs to

- estimate error decay  $|\mathbb{E}[Q - Q_\ell]|$ :  $\rightsquigarrow$  needed to determine optimal  $L$
- estimate variance decay  $V_\ell$ :  $\rightsquigarrow$  needed to determine optimal  $\{M_\ell\}_{\ell=0}^L$

$|\mathbb{E}[Q - Q_\ell]|$  can be estimated as  $|\mu_\ell^{MC} - \mu_{\ell-1}^{MC}|$  based on a pilot run

$V_\ell$  can be estimated by sample variance estimator based on pilot runs

**Problem:** on the finest levels we should run only very few simulations.  
 Cost for estimation of  $V_L$  might dominate the overall cost of the MLMC algorithm.

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# Continuation Multilevel Monte Carlo

[Collier-HajiAli-N.-vonSchwerin-Tempone 2015, Pisoni-N.-Leyland 2017]

**Idea:** Solve the problem with decreasing tolerances  $tol^{(0)} > tol^{(1)} > \dots \geq tol$ .  
Use collected samples on all levels to improve the estimate of  $V_\ell$  and  $|\mathbb{E}[Q - Q_\ell]|$ .

Estimator  $\hat{V}_\ell$  of  $V_\ell = \text{Var}[\Delta Q_\ell]$  at iteration  $j$ : MAP Bayesian estimator

- we make the ansatz  $\Delta Q_\ell \sim N(\mu_\ell, V_\ell)$
- based on acquired samples at previous iteration, we fit models (least squares)
  - $\mu_\ell^{model} = c_\alpha h_\ell^\alpha$
  - $V_\ell^{model} = c_\beta h_\ell^\beta$
- We take a Normal-Gamma prior for  $(\mu_\ell, V_\ell)$ , with mode in  $(\mu_\ell^{model}, V_\ell^{model})$
- Then  $\hat{V}_\ell$  is the MAP Bayesian estimator based on the Normal-Gamma prior and the actual samples acquired at iteration  $j$

Effectively, we have

$$\begin{array}{lll}
 M_\ell = 0 & \hat{V}_\ell = V_\ell^{model} & \text{(prior model)} \\
 M_\ell \rightarrow \infty & \hat{V}_\ell \approx V_\ell^{MC} & \text{(sample variance)}
 \end{array}$$

$\hat{V}_\ell$  is then used to determine the sample sizes  $M_\ell$  for the next iteration.



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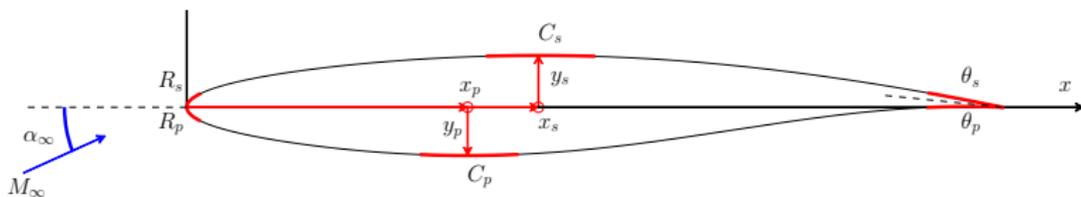
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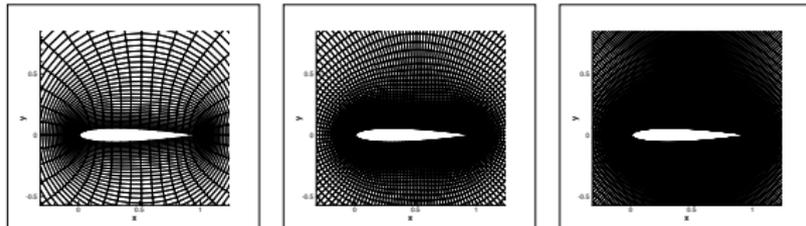


# Computation of pressure coefficient for NACA 0012 / NASA SC(2)-0012 airfoils

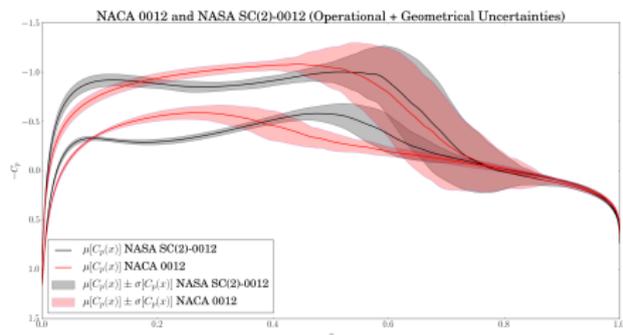
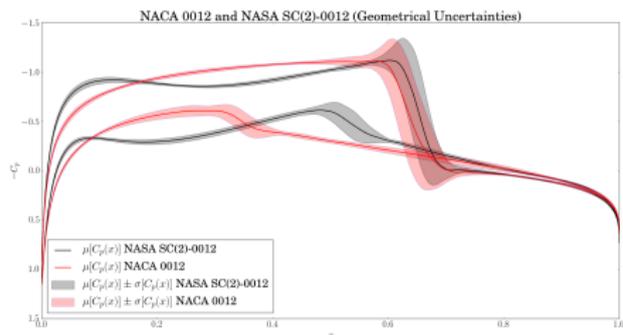
	Name	Nominal value	Uncertainty
<b>Operational</b>	$T_\infty$	$T_n = 288.15 [K]$	$\mathcal{TN}(T_n, 2\%, 110\%, 90\%)$
	$p_\infty$	$p_n = 101325 [N/m^2]$	$\mathcal{TN}(p_n, 2\%, 110\%, 90\%)$
	$\alpha$	$\alpha_n = 1.25^\circ$	$\mathcal{TN}(\alpha_n, 1\%, 110\%, 90\%)$
	M	$M_n = 0.8$	$\mathcal{TN}(M_n, 2\%, 110\%, 90\%)$
<b>Geometrical</b>	$R_p$	0.01458398	$\mathcal{TN}(R_{P_n}, 2.5\%, 110\%, 90\%)$
	$R_s$	0.01458398	$\mathcal{TN}(R_{S_n}, 2.5\%, 110\%, 90\%)$
	$X_p$	0.30049047	$\mathcal{TN}(X_{P_n}, 2.5\%, 110\%, 90\%)$
	$X_s$	0.30049047	$\mathcal{TN}(X_{S_n}, 2.5\%, 110\%, 90\%)$
	$Y_p$	-0.05994286	$\mathcal{TN}(Y_{P_n}, 2.5\%, 110\%, 90\%)$
	$Y_s$	0.05994286	$\mathcal{TN}(Y_{S_n}, 2.5\%, 110\%, 90\%)$
	$C_p$	0.44213792	$\mathcal{TN}(C_{P_n}, 2.5\%, 110\%, 90\%)$
	$C_s$	-0.44213792	$\mathcal{TN}(C_{S_n}, 2.5\%, 110\%, 90\%)$
	$\theta_p$	8.3763395	$\mathcal{TN}(\theta_{P_n}, 2.5\%, 110\%, 90\%)$
	$\theta_s$	-8.3763395	$\mathcal{TN}(\theta_{S_n}, 2.5\%, 110\%, 90\%)$



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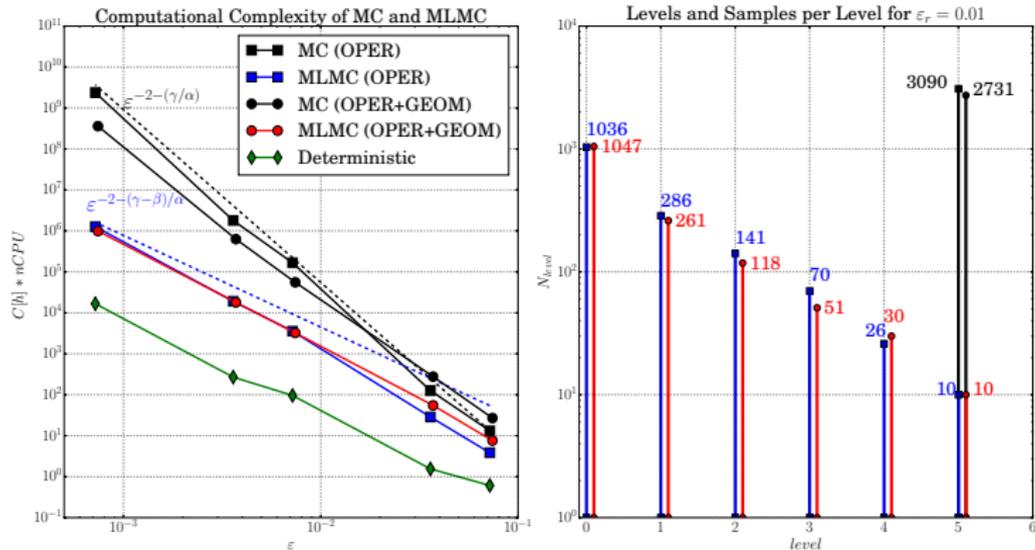


LEVEL	Airfoil nodes	Cells	Avg. Real Computational Time [s] (CPU)
L0	41	6943	12.4 (32)
L1	81	11115	20.9 (38)
L2	161	19385	26.9 (44)
L3	321	36251	71.1 (50)
L4	641	71477	231.15 (56)
L5	1281	145005	422.0 (64)

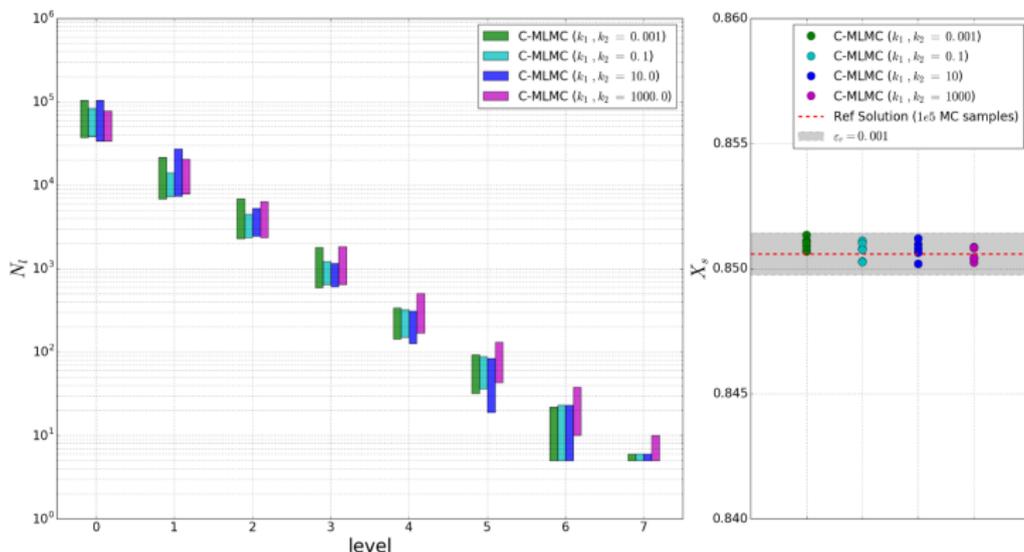


Inviscid model (Euler); SU2 solver (Stanford) [Pisaroni-Leyland-N., AIAA Aviation, 2016]

## MLMC vs MC for aerodynamic inviscid problems



# Robustness of C-MLMC estimator



Variability over 10 repetitions of the C-MLMC algorithm for different parameters in the Normal-Gamma prior.

# Outline

- 1 Motivating example
- 2 Multilevel Monte Carlo for expectations
- 3 MLMC for moments and distributions**
- 4 Risk averse optimization with MLMC
- 5 Conclusions

# Beyond expectations: computation of central moments

**Goal:** compute  $\mu_p(Q) = \mathbb{E}[(Q - \mathbb{E}[Q])^p]$

How to apply and tune MLMC in this case? [Bierig-Chernov 2015-2016] use biased central moments estimators.

Alternatively, use *h-statistics* [Pisaroni-Krumscheid-N. 2017]. Given iid sample  $\vec{Q}_M = \{Q^{(1)}, \dots, Q^{(M)}\}$ ,

$h_p(\vec{Q}_M)$ : unbiased estimator of  $\mu_p(Q)$  with minimal variance

**Multilevel estimator:** 
$$h_p^{MLMC} = \sum_{\ell=0}^L (h_p(\vec{Q}_{\ell, M_\ell}) - h_p(\vec{Q}_{\ell-1, M_\ell}))$$

with  $(\vec{Q}_{\ell, M_\ell}, \vec{Q}_{\ell-1, M_\ell})$  generated with the same noise (highly correlated)

**Mean squared error:** 
$$\text{MSE}(h_p^{MLMC}) = (\mu_p(Q) - \mu_p(Q_L))^2 + \sum_{\ell=0}^L \frac{V_{\ell, p}}{M_\ell}$$

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**Complexity result** for  $h_\ell = h_0 s^{-\ell}$

Assume  $\mu_{2p}(Q_\ell) < \infty$  for all  $\ell$  and there exist  $\alpha, \beta, \gamma > 0$ ,  $2\alpha \geq \min\{\beta, \gamma\}$  s.t.

- $|\mu_p(Q) - \mu_p(Q_\ell)| = \mathcal{O}(h_\ell^\alpha)$ ,
- $V_{\ell,p} = \mathcal{O}(h_\ell^\beta)$ ,
- $C_\ell = \text{Cost}(Q_\ell^{(i,\ell)}, Q_{\ell-1}^{(i,\ell)}) = \mathcal{O}(h^{-\gamma})$ ,

Then, taking  $L = \mathcal{O}(\text{tol}^{\frac{1}{\alpha}})$  and  $M_\ell = \left\lceil \text{tol}^{-2} \sqrt{\frac{V_{\ell,p}}{C_\ell}} \left( \sum_{k=0}^L \sqrt{C_k V_{k,p}} \right) \right\rceil$  leads to

$$\text{MSE}(h_p^{\text{MLMC}}) \lesssim \text{tol}^2 \quad \text{and} \quad W(h_p^{\text{MLMC}}) \lesssim \begin{cases} \text{tol}^{-2}, & \beta > \gamma \\ \text{tol}^{-2} |\log(\text{tol})|^2, & \beta = \gamma \\ \text{tol}^{-2 - \frac{\gamma - \beta}{\alpha}}, & \beta < \gamma \end{cases}$$

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**Technical difficulty:** how to estimate the variances  $V_{\ell,p}$

Define  $\vec{X}_{\ell,M_\ell}^+ = \vec{Q}_{\ell,M_\ell} + \vec{Q}_{\ell-1,M_\ell}$ ,  $\vec{X}_{\ell,M_\ell}^- = \vec{Q}_{\ell,M_\ell} - \vec{Q}_{\ell-1,M_\ell}$

$\Delta_\ell h_p = h_p(\vec{Q}_{\ell,M_\ell}) - h_p(\vec{Q}_{\ell-1,M_\ell})$  can be expressed as a power sum

$$\Delta_\ell h_p = \sum_{a+b \leq p} S_{a,b}(\vec{X}_{\ell,M_\ell}^+, \vec{X}_{\ell,M_\ell}^-), \quad S_{a,b}(\vec{X}, \vec{Y}) = \sum_i (X^{(i)})^a (Y^{(i)})^b$$

Unbiased estimators  $\hat{V}_{\ell,p}$  of  $V_{\ell,p}$  can be computed in closed form starting from the power terms  $S_{a,b}(\vec{X}_{\ell,M_\ell}^+, \vec{X}_{\ell,M_\ell}^-)$  [Pisaroni-Krumscheid-N. 2017].

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# Beyond expectations: char. function, CDF, and more

Some derived quantities can be written as parametric expectations

**Example 1:** Characteristic function of  $Q$

$$\Phi(\theta) = \mathbb{E}[\phi(\theta, Q)], \quad \phi(\theta, Q) = e^{i\theta Q}$$

$\rightsquigarrow$  we can compute  $\Phi(\theta_j)$  by MLMC on a set of points  $\theta_j$ .

**Example 2:** CDF of  $Q$

$$F(\theta) = \mathbb{E}[\phi(\theta, Q)], \quad \phi(\theta, Q) = \mathbb{1}_{\{Q \leq \theta\}}$$

**Problem:**  $\phi(\theta, Q)$  is not smooth ! When applying MLMC, the variance of the differences,  $V_\ell = \text{Var}[\phi(\theta, Q_\ell) - \phi(\theta, Q_{\ell-1})]$  will decay slowly. **Not much gain in MLMC.**

**Remedies:**

- [Giles-Nagapetyan-Ritter 2015] smoothing:  $F_\epsilon(\theta) = \mathbb{E}[\phi_\epsilon(\theta, Q)]$ . Technical difficulty:  $\epsilon$  should depend on the required tolerance  $\rightsquigarrow$  difficult tuning of MLMC
- [Bierig-Chernov 2016] approximate  $F$  or pdf based on moments (see Alexey's talk)
- [Krumscheid-N. 2017] anti-derivative approach:  $F(\theta) = \Phi'(\theta)$  with  $\Phi(\theta) = \mathbb{E}[\phi(\theta, Q)]$  and  $\phi(\theta, \cdot)$  Lipschitz continuous.



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**Problem:**  $\phi(\theta, Q)$  is not smooth ! When applying MLMC, the variance of the differences,  $V_\ell = \text{Var}[\phi(\theta, Q_\ell) - \phi(\theta, Q_{\ell-1})]$  will decay slowly. **Not much gain in MLMC.**

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- [Giles-Nagapetyan-Ritter 2015] smoothing:  $F_\epsilon(\theta) = \mathbb{E}[\phi_\epsilon(\theta, Q)]$ . Technical difficulty:  $\epsilon$  should depend on the required tolerance  $\rightsquigarrow$  difficult tuning of MLMC
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# Beyond expectations: char. function, CDF, and more

Some derived quantities can be written as parametric expectations

**Example 1:** Characteristic function of  $Q$

$$\Phi(\theta) = \mathbb{E}[\phi(\theta, Q)], \quad \phi(\theta, Q) = e^{i\theta Q}$$

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For any  $\tau \in (0, 1)$  define

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Then

$$F(\theta) = (1 - \tau)\Phi'_\tau(\theta) + \tau$$

and MLMC can be effectively used to approximate  $\Phi_\tau(\theta)$  and its derivatives.

Moreover, from the approximation of  $\Phi_\tau$  and its derivatives we can get for free

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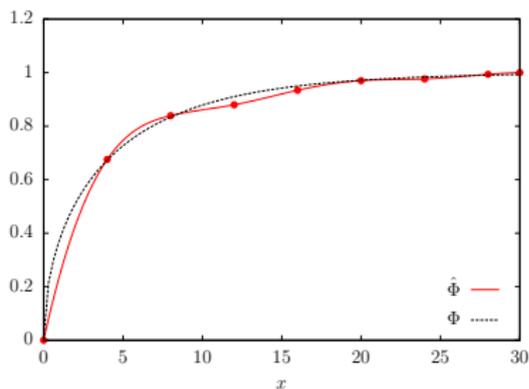
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# Computing parametric expectations by MLMC

**Goal:** given  $\phi(\theta, Q)$ , approximate  $\Phi(\theta) = \mathbb{E}[\phi(\theta, Q)]$  and its derivatives **uniformly** in  $\Theta$ .



Interpolation approach:

- introduce a grid  $\vec{\xi} = \{\xi_1, \dots, \xi_n\} \subset \Theta$
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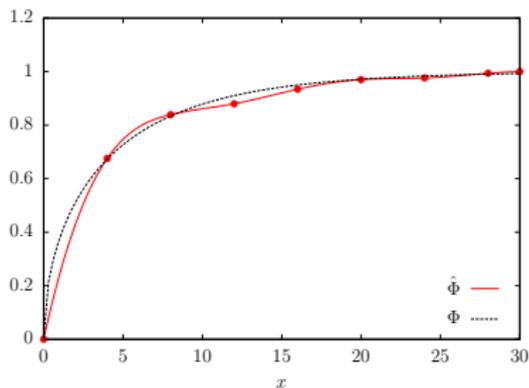
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Assumptions on  $\mathcal{I}_n$  (valid for spline interpolation)

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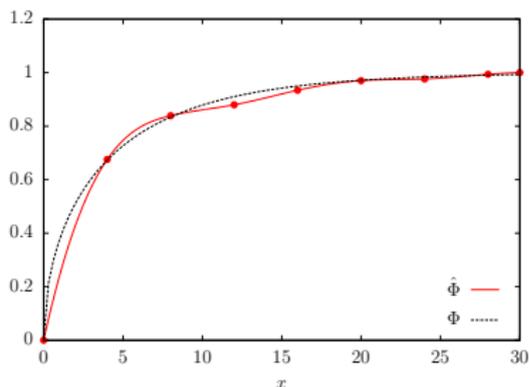
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# Error splitting

Define the **mean squared error**:  $\text{MSE}(\hat{\Phi}_L) = \mathbb{E}[\|\Phi - \hat{\Phi}_L\|_{L^\infty(\Theta)}^2]$

**Notation**: for  $\vec{x} \in \mathbb{R}^n$  define  $\text{Var}[\vec{x}] = \mathbb{E}[\|\vec{x} - \mathbb{E}[\vec{x}]\|_{\ell^\infty}^2]$

**Useful result**: for  $\vec{x}^{(1)}, \dots, \vec{x}^{(k)} \in \mathbb{R}^n$  independent,

$$\text{Var}\left[\sum_{i=1}^k \vec{x}^{(i)}\right] \leq c \log(n) \sum_{i=1}^k \text{Var}[\vec{x}^{(i)}]$$

## Error splitting

$$\begin{aligned} \text{MSE}(\hat{\Phi}_L) &\leq 3\|\Phi - \mathcal{I}_n\Phi\|_\infty^2 + 3\|\mathcal{I}_n\Phi - \mathcal{I}_n\Phi_L\|_\infty^2 + 3\mathbb{E}[\|\mathcal{I}_n\Phi_L - \mathcal{I}_n\Phi_L^{\text{MLMC}}\|_\infty^2] \\ &\lesssim \underbrace{\|\Phi - \mathcal{I}_n\Phi(\vec{\xi})\|_\infty^2}_{\text{interp. error}} + \underbrace{\|\Phi(\vec{\xi}) - \Phi_L(\vec{\xi})\|_\infty^2}_{\text{discret. error}} + \log(n) \underbrace{\sum_{\ell=0}^L \frac{V_\ell}{M_\ell}}_{\text{statistical error}} \end{aligned}$$

with  $V_\ell = \text{Var}[\phi(\vec{\xi}, Q_\ell) - \phi(\vec{\xi}, Q_{\ell-1})]$ . All terms can be estimated in practice.  
Optimization of MLMC based on estimators  $\hat{V}_\ell$ .

[Pisaroni-Krumscheid-N. in preparation]

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# Complexity analysis

**Complexity result** for  $h_\ell = h_0 5^{-\ell}$  [Krumscheid-N. 2017]

Assume

- $\|\Phi - \Phi_\ell\|_{L^\infty(\Theta)} \leq c_1 h_\ell^\alpha$ ,
- $\mathbb{E}\left(\|\phi(\cdot, Q_\ell) - \phi(\cdot, Q_{\ell-1})\|_{L^\infty(\Theta)}^2\right) \leq c_2 h_\ell^\beta$ ,
- cost to simulate one realization of  $\phi(\theta, Q_\ell) \leq c_3 h_\ell^{-\gamma}$ .

If  $\Phi \in C^{k+1}(\Theta)$ , there exists an estimator  $\hat{\Phi}_L$  s.t.  $\text{MSE}(\hat{\Phi}_L) = \mathcal{O}(\text{tol}^2)$  and

$$W(\hat{\Phi}_L) \lesssim \text{tol}^{-(2+\frac{1}{k+1})} |\log(\text{tol})| + |\log(\text{tol})| \begin{cases} \text{tol}^{-2}, & \text{if } \beta > \gamma, \\ \text{tol}^{-2} |\log(\text{tol})|^2, & \text{if } \beta = \gamma, \\ \text{tol}^{-(2+\frac{\gamma-\beta}{\alpha})}, & \text{if } \beta < \gamma, \end{cases}$$

The first term accounts for the cost of computing the spline interpolation. This is often negligible for heavy computational models. It can be removed by taking  $n = n_\ell$  (different spline interpolant on each level).

Neglecting the first term, the complexity is essentially the same as for simple expectations, up to an extra log factor.

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## Complexity result for derivatives [Krumscheid-N. 2017]

If  $\Phi \in C^{2k+2}(\Theta)$  and  $m \leq 2k + 1$ , there exists an estimator  $\hat{\Phi}_L$  s.t.  
 $\mathbb{E}[\|\frac{d^m}{d\theta^m} \Phi - \frac{d^m}{d\theta^m} \hat{\Phi}_L\|_\infty^2] = \mathcal{O}(tol^2)$  and

$$W(\hat{\Phi}_L) \lesssim |\log(tol)| \begin{cases} tol^{-2\frac{2k+2}{2k+2-m}}, & \text{if } \beta > \gamma, \\ tol^{-2\frac{2k+2}{2k+2-m}} |\log(tol)|^2, & \text{if } \beta = \gamma, \\ tol^{-(2+\frac{\gamma-\beta}{\alpha})\frac{2k+2}{2k+2-m}}, & \text{if } \beta < \gamma, \end{cases}$$

(neglecting the cost of interpolation)

This result applies to the approximation of CDF, quantiles and CVaR with  $m = 1$  and PDF with  $m = 2$ .

## An example: the characteristic function

- An SDE model to describe a **European call option**, where the asset follows

$$dS = rS dt + \sigma S dW, \quad S(0) = S_0,$$

- **Quantity of interest** is the discounted “payoff”:  $Q := e^{-rT} \max(S(T) - K, 0)$
- Approximate **characteristic function** of  $Q$ :

$$\Phi(\theta) = \mathbb{E}(\cos(\theta Q)) + i \mathbb{E}(\sin(\theta Q)) \equiv \Phi_1(\theta) + i \Phi_2(\theta),$$

- Milstein scheme with  $h_\ell = 2^{-\ell} T$ ;  $\Theta = [-1, 1]$ ,  $r = \frac{1}{20}$ ,  $\sigma = \frac{1}{5}$ ,  $T = 1$ ,  $K = 10 = S_0$ .

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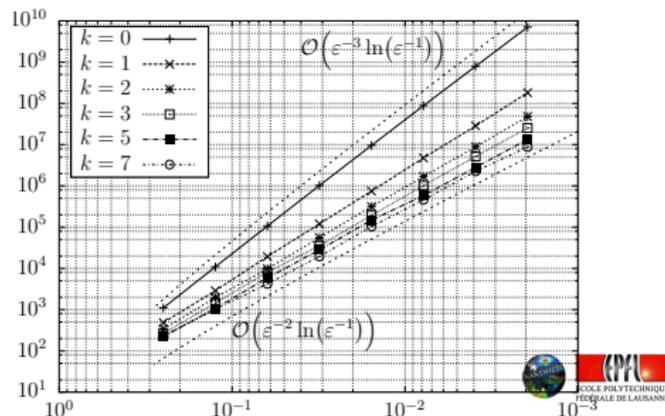
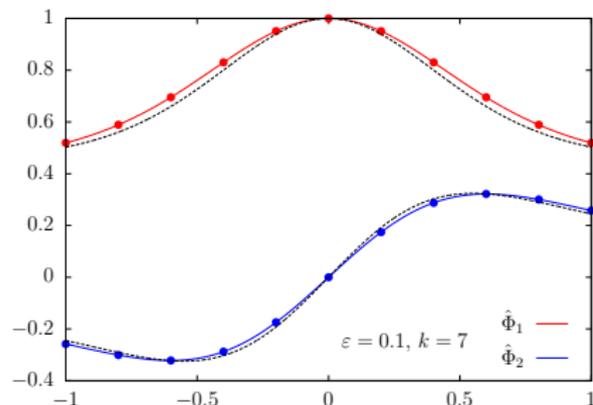
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# NASA Common Research Model

NASA CRM: aircraft configuration equipped with a contemporary supercritical transonic wing and a fuselage that is representative of a wide-body commercial transport aircraft.

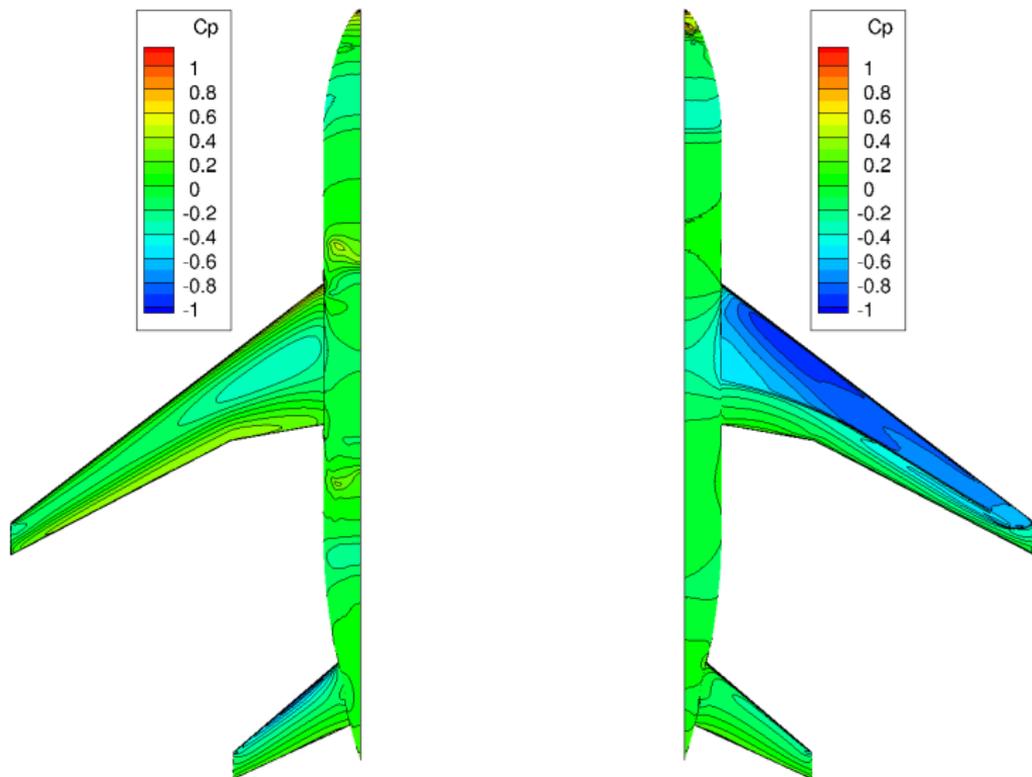
Q.ty	Reference	Uncertainty
$M_\infty$	0.85	$\mathcal{B}(2, 2, 0.05, M_\infty - 0.025)$
$Re_c$	$5 \cdot 10^6$	—
$T_{ref}$	310.928 [K]	$\mathcal{B}(2, 2, 30, T_{ref} - 15)$
$C_L$	0.3, 0.4, 0.5, 0.55	—



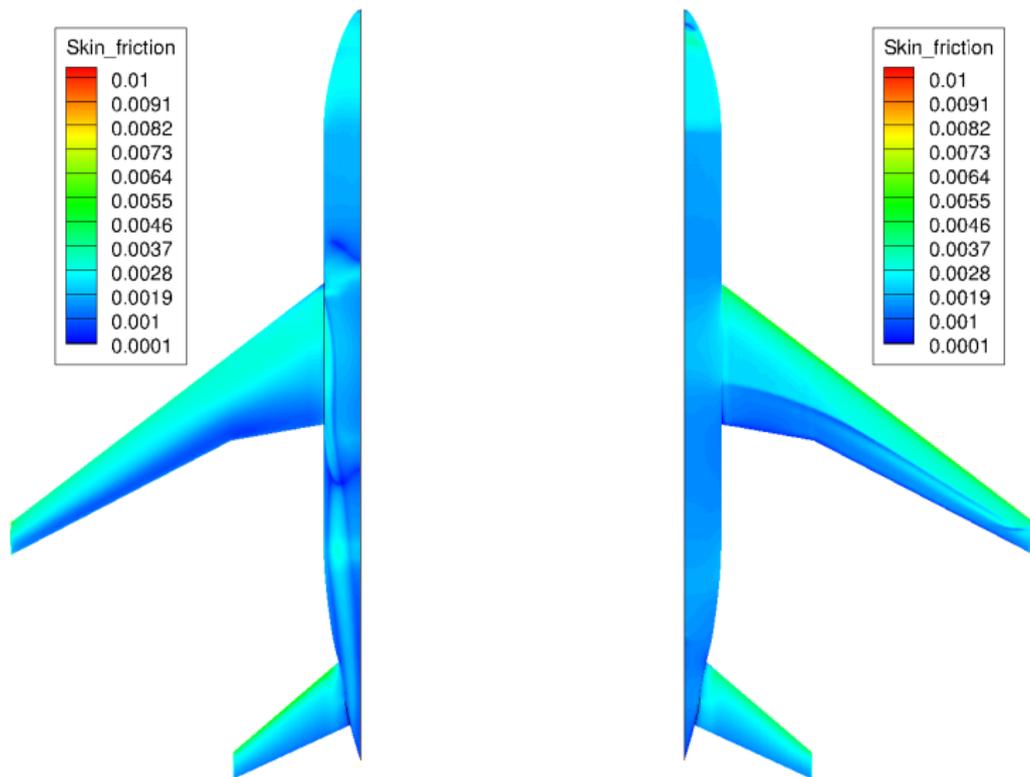
LEVEL	Cells	y+	CTime on 280 CPUs
L0	$2.3 \cdot 10^6$	1 – 2	400 [s] (0.11 [h])
L1	$5.0 \cdot 10^6$	1 – 2	825 [s] (0.23 [h])
L2	$9.8 \cdot 10^6$	1 – 2	1250 [s] (0.35 [h])
L3	$21.3 \cdot 10^6$	1 – 2	3200 [s] (0.89 [h])

Spalart-Allmaras turbulence model, hybrid unstructured grids.

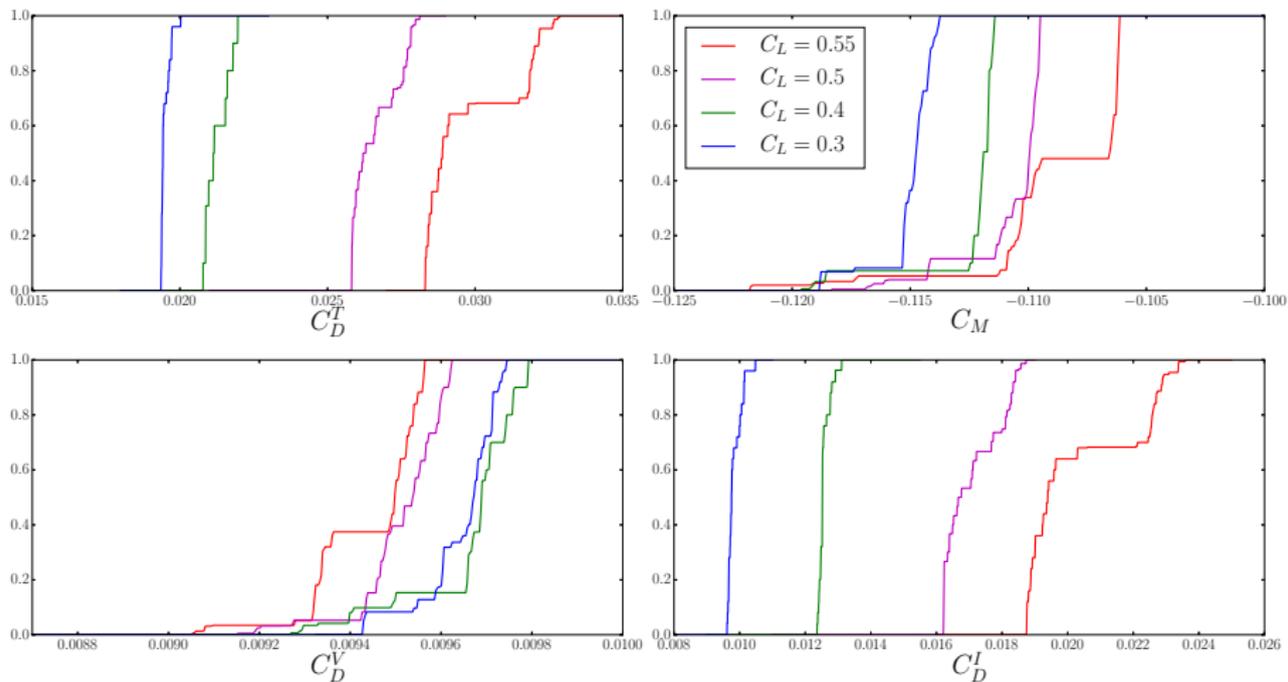
# NASA Common Research Model



# NASA Common Research Model



# NASA Common Research Model



# Outline

- 1 Motivating example
- 2 Multilevel Monte Carlo for expectations
- 3 MLMC for moments and distributions
- 4 Risk averse optimization with MLMC**
- 5 Conclusions

# Risk averse optimization

$$\min_{x \in X} \mathcal{R}(Q(x)), \quad X: \text{feasible design space}$$

$\mathcal{R}$ : risk measure

## Examples

- $\mathcal{R}(Q) = \mathbb{E}[Q]$  (mean-based risk)
- $\mathcal{R}(Q) = \mathbb{E}[Q] \pm \alpha \text{std}[Q]$
- $\mathcal{R}(Q) = q_\alpha[Q]$  ( $\alpha$ -quantile)
- $\mathcal{R}(Q) = \text{CVaR}_\alpha[Q]$

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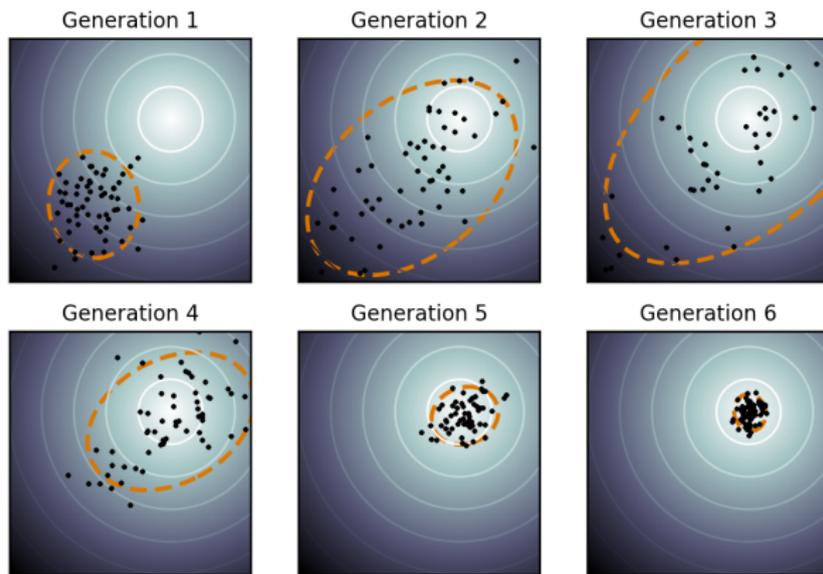
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# Combining MLMC with CMA-ES

Optimization done by Covariance Matrix Adaptation Evolutionary Algorithm (CMA-ES)



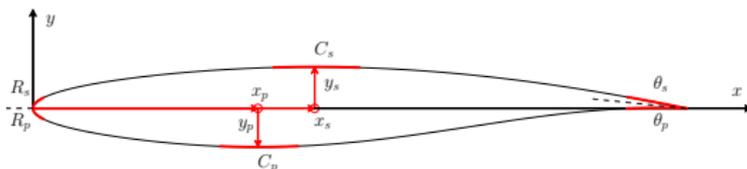
For each individual at each generation, risk measure computed by MLMC.

## Airfoil optimization under operating uncertainties

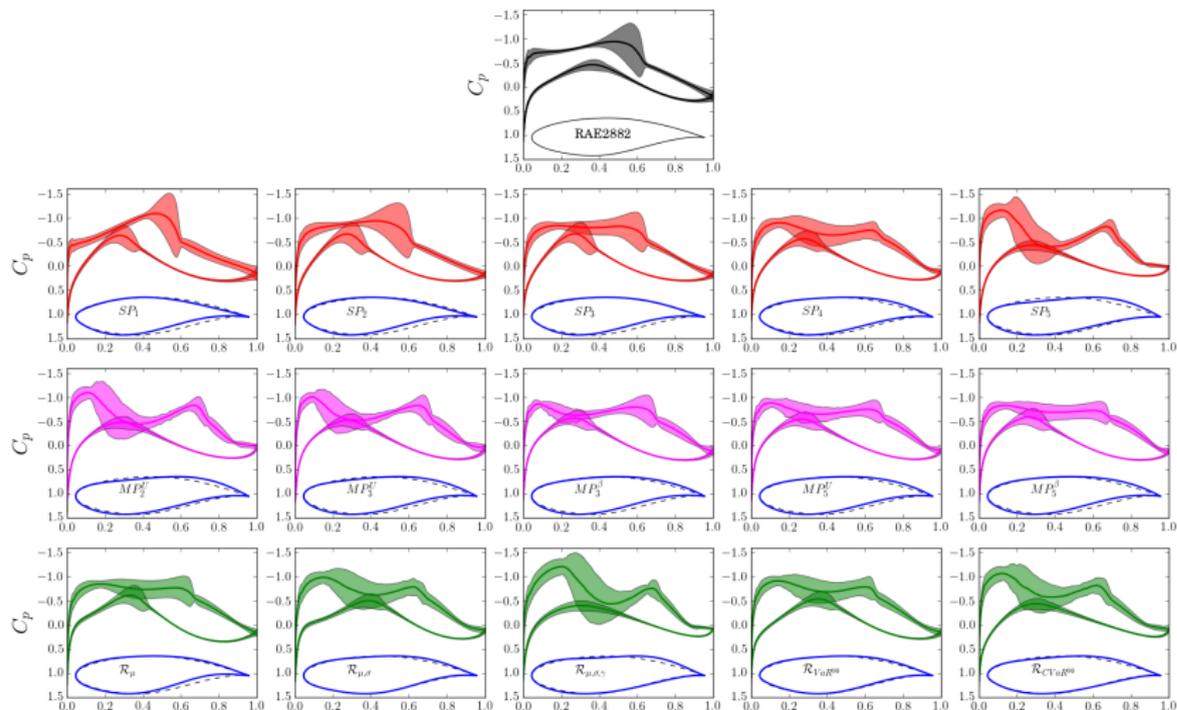
$$\begin{cases} \min_{x \in X} \mathcal{R}[C_D(x)] \\ \text{s.t. } C_L(x) = C_L^*, \quad \text{thickness constraint} \end{cases}$$

$\mathcal{R}_{\mu, \sigma}[C_D(x)]$	$\mu_{C_D}(x) + \sigma_{C_D}(x)$
$\mathcal{R}_{\mu, \sigma, \gamma}[C_D(x)]$	$\mu_{C_D}(x) + \sigma_{C_D}(x) + \mu_{C_D}(x) \cdot \gamma_{C_D}(x)$
$\mathcal{R}_{VaR^{90}}[C_D(x)]$	$VaR_{C_D}^{90}(x)$
$\mathcal{R}_{CVaR^{90}}[C_D(x)]$	$CVaR_{C_D}^{90}(x)$

	Quantity	Reference ( $r$ )	Uncertainty
Operating parameters	$C_L$	0.5	—
	$M_\infty$	0.75	$\mathcal{B}(2, 2, 0.1, M_\infty - 0.05)$
	$Re_c$	$6.5 \cdot 10^6$	—
	$p_\infty$ [Pa]	101325	—
	$T_\infty$ [K]	288.5	—

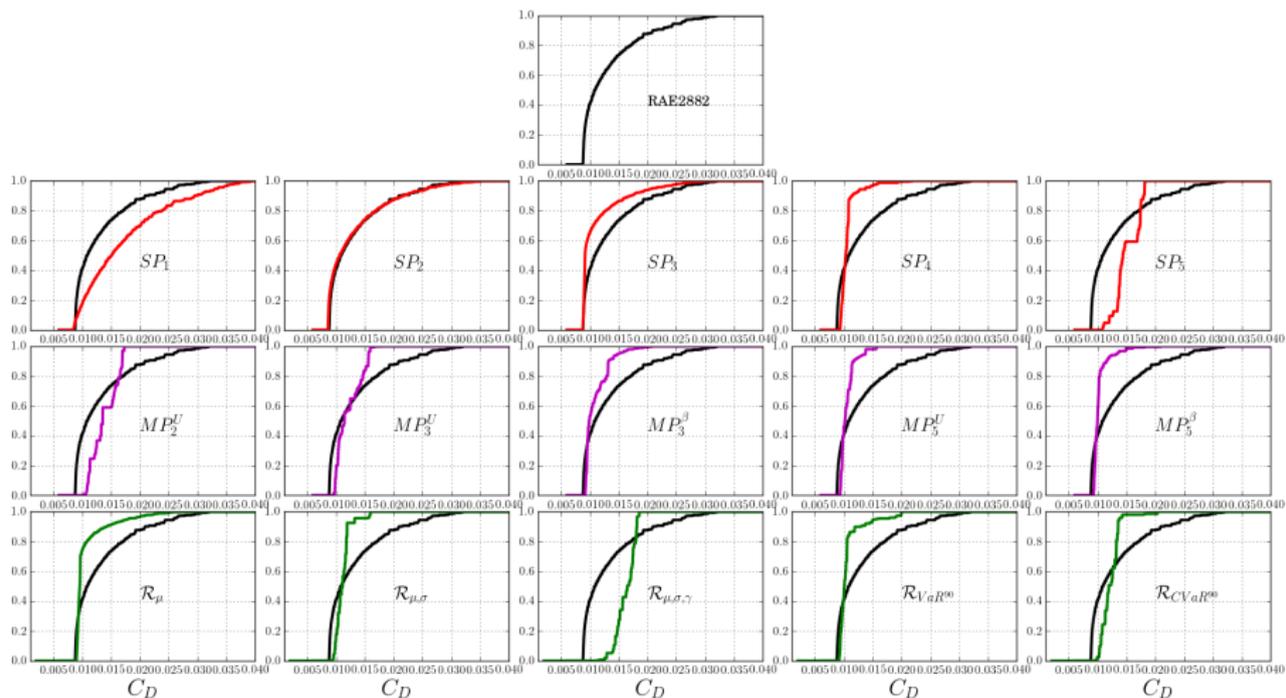


# Qualitative comparison



Model: steady state Euler + boundary layer equation (MSES software)

# Deterministic versus Robust optimization

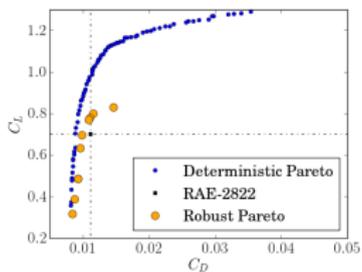
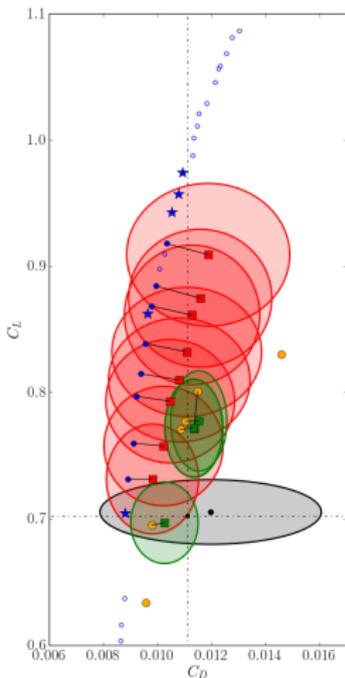
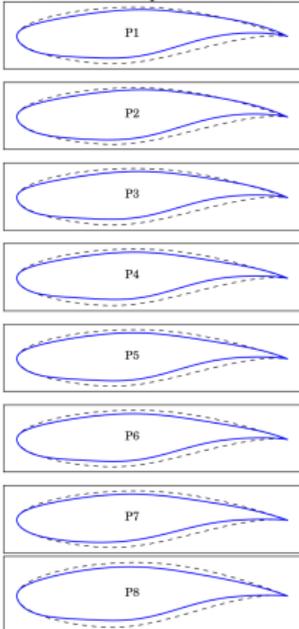


# Multi-objective optimization under operating uncertainties

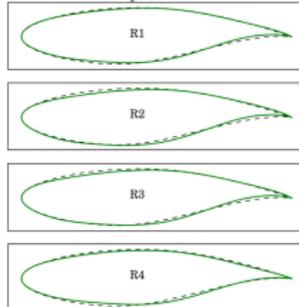
$$P\text{-min}_{x \in X} \{ \mu_{C_D}(x) + \sigma_{C_D}(x), -\mu_{C_L}(x) + \sigma_{C_L}(x) \} \quad (\text{Pareto front})$$

Uncertainties in Mach number and Angle of Attack.

Deterministic Optimized Airfoils



Robust Optimized Airfoils



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# Conclusions and outlook

- Multilevel Monte Carlo is a very powerful technique that can dramatically reduce the computational cost of a UQ analysis compared to plain MC.
- The tuning of MLMC requires adaptive algorithms and reliable error and variances estimators.
- We have presented a way to compute higher order moments as well as cdf, quantiles, CVaR with MLMC and properly tune the method.
- The methodology has been successfully applied to forward UQ propagation and robust optimization under uncertainty in compressible aerodynamics.

Thank you for your attention!

# References



M. Pisaroni.

Multi Level Monte Carlo Methods for Uncertainty Quantification and Robust Design Optimization in Aerodynamics, [PhD Thesis n. 8082, EPFL, 2017](#)



M. Pisaroni, S. Krumscheid, F. Nobile.

Quantifying uncertain system outputs via the multilevel Monte Carlo method - Part I: Central moment estimation, [MATHICSE Technical report no. 23.2017](#).



M. Pisaroni, S. Krumscheid, F. Nobile.

Quantifying uncertain system outputs via the multilevel Monte Carlo method Part 2: distribution and robustness measures, [in preparation](#).



S. Krumscheid, F. Nobile.

Multilevel Monte Carlo approximation of functions, [MATHICSE Technical report no. 12.2017](#).



M. Pisaroni, F. Nobile, P. Leyland.

A Multilevel Monte Carlo Evolutionary Algorithm for Robust Aerodynamic Shape Design. [18th AIAA/ISSMO Multidisciplinary Analysis and Optimization Conference, Denver, Colorado, USA, 2017](#).



M. Pisaroni, F. Nobile, P. Leyland.

Continuation Multi-Level Monte-Carlo method for Uncertainty Quantification in Turbulent Compressible Aerodynamics Problems modeled by RANS, [MATHICSE Technical report no. 10.2017](#).



M. Pisaroni, F. Nobile, P. Leyland.

A Continuation Multi Level Monte Carlo (C-MLMC) method for uncertainty quantification in compressible inviscid aerodynamics, [CMAME, vol. 326, p. 20-50, 2017](#).



A.-L. Haji-Ali, F. Nobile, R. Tempone.

Multi-index Monte Carlo: when sparsity meets sampling, [in Numer. Math., vol. 132\(4\), p. 767-806, 2016](#).



N. Collier, A.-L. Haji-Ali, F. Nobile, E. von Schwerin, R. Tempone.

A continuation multilevel Monte Carlo algorithm, [BIT Numerical Mathematics, vol. 55\(2\), p. 399-432, 2015](#).